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A Supplement for D. P. No. 713

by

Yoshiko Nogami

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Summary.

This paper is added as a supplement for D. P. No. 713.

This paper is added as a supplement to the proof of UMP-unbiasedness for $\phi^*(y)$ given by (5) in D.P. No. 713.

Statement: Let X_1, \dots, X_n be a random sample from p.d.f. $f(x|\theta)=1$ for $\theta \leq x < \theta+1$; $=0$, otherwise. The test ϕ^* given by (5) in D. P. No. 713 for testing $H_0: \theta = \theta_0$ versus $H_1: \theta \neq \theta_0$ is UMP-unbiased size- α .

Proof.) Let $Y = (X_{(1)} + X_{(n)} - 1)/2$. The p.d.f. of Y is given by

$$(1) \quad g(y|\theta) = n(1-2|y-\theta|)^{n-1} I_{(-1/2, 1/2)}(y-\theta) \quad (\dot{=} g(y-\theta)).$$

Let ϕ be a size- α test symmetric about θ . Let $x_\alpha(\theta) = E_\theta(\phi(Y))$ and we abbreviate $\phi_{\theta_0}(y)$ as $\phi(y)$. Then, it follows that for given α such that

$$0 < \alpha < 1,$$

$$(2) \quad x_\alpha(\theta_0) = E_{\theta_0}(\phi(Y)) = \alpha$$

and

$$(3) \quad E_{\theta_0}(Y\phi(Y)) = \theta_0 E_{\theta_0}(\phi(Y)) = \theta_0 \alpha.$$

Above (3) holds because $E_{\theta_0}((Y-\theta_0)\phi(Y)) = 0$ and (2) holds.

Hence, by generalized fundamental lemma, ϕ maximizes the integral

$$(4) \quad \int \phi(y)g(y|\theta')dy \quad \text{for } \theta' \neq \theta_0$$

when there exist k_1 and k_2 such that

$$(5) \quad \phi(y) = \begin{cases} 1, & \text{if } (1-2|y-\theta'|)^{n-1} I_{(-1/2, 1/2)}(y-\theta') \\ & \geq k_2(1-2|y-\theta_0|)^{n-1} I_{(-1/2, 1/2)}(y-\theta_0) \\ & \quad + k_1 y(1-2|y-\theta_0|)^{n-1} I_{(-1/2, 1/2)}(y-\theta_0), \\ 0, & \text{if } (1-2|y-\theta'|)^{n-1} I_{(-1/2, 1/2)}(y-\theta') \\ & < k_2(1-2|y-\theta_0|)^{n-1} I_{(-1/2, 1/2)}(y-\theta_0) \\ & \quad + k_1 y(1-2|y-\theta_0|)^{n-1} I_{(-1/2, 1/2)}(y-\theta_0). \end{cases}$$

In order for $x_\alpha(\theta)$ to be minimized at $\theta = \theta_0$, the region of y must have two-sided intervals.

Let $h(y) = (1 - 2|y - \theta'|) / (1 - 2|y - \theta_0|)$ (> 0). The first inequality in (5)

(we denote this by FI(5)) becomes

$$(6) \quad h(y) \geq (k_2 + k_1 y)^{1/(n-1)} (\doteq z(y))$$

$$(7) \quad \text{for } \theta_0 < \theta' \quad \& \quad \theta' - 1/2 < y < \theta_0 + 1/2, \quad \text{or for } \theta' < \theta_0 \quad \& \quad \theta_0 - 1/2 < y < \theta' + 1/2.$$

Let $\theta_0 < \theta'$. Then, for $\theta_0 - 1/2 < y < \theta' - 1/2$, FI(5) holds when $k_1 = k_2 = 0$, say. On the other hand, for $\theta_0 + 1/2 < y < \theta' + 1/2$, FI(5) always holds for any real k_1 and k_2 . Let $\theta' < \theta_0$. Then, for $\theta' - 1/2 < y < \theta_0 - 1/2$, FI(5) always holds for any real k_1 and k_2 . On the other hand, for $\theta' + 1/2 < y < \theta_0 + 1/2$, FI(5) holds when $k_1 = k_2 = 0$, say.

Since $(0 <) h(\theta_0) = 1 - 2|\theta_0 - \theta'| (< 1)$, the graph of $h(y)$ becomes as solid lines in the figure, below. We now examine (6). In order to get two-sided intervals for y according to (6), we take $k_1 > 0$ for $\theta_0 < \theta'$ and $k_1 < 0$ for $\theta' < \theta_0$ and for simplicity we assume $z(\theta' - 1/2) = 0$ for $\theta_0 < \theta'$ and $z(\theta' + 1/2) = 0$ for $\theta' < \theta_0$. Then, k_2 is taken as follows:

$$(8) \quad k_2 = \begin{cases} -k_1(\theta' - 1/2), & \text{for } \theta_0 < \theta', \\ -k_1(\theta' + 1/2), & \text{for } \theta' < \theta_0. \end{cases}$$

Substituting these values into $z(y)$ we have

$$(9) \quad z(y) = \begin{cases} (k_1(y - \theta' + 1/2))^{1/(n-1)}, & \text{for } \theta_0 < \theta', \\ (k_1(y - \theta' - 1/2))^{1/(n-1)}, & \text{for } \theta' < \theta_0, \end{cases}$$

which is drawn by stripe lines in the figure, below.

Since we need $h(\theta_0) < z(\theta_0)$, we take $k_1 = 2$ for $\theta_0 < \theta'$ and $k_1 = -2$ for $\theta' < \theta_0$.

Then, $z^{n-1}(\theta_0) = h(\theta_0)$, for all n and hence, there exists an integer n_0 such that for all $n > n_0$, $h(\theta_0) < z(\theta_0) (< 1)$.

Therefore, if we denote by y_1 and y_2 , y -coordinates of the points of intersection of $z(y)$ and $h(y)$ for $\theta' < \theta_0$ and for $\theta' > \theta_0$, respectively, then our two-sided intervals are given as follows:

$$(10) \quad y < \theta' - 1/2 \quad \text{or} \quad y > y_2 \quad , \text{for } \theta_0 < \theta',$$

$$(11) \quad y < y_1 \quad \text{or} \quad y > \theta' + 1/2 \quad , \text{for } \theta' < \theta_0.$$

Now, we let $y_0 = -k_2/k_1$. For $\theta_0 < \theta'$, take $\theta' - 1/2 < y_0 < \theta_0$ and for $\theta' < \theta_0$, take $\theta_0 < y_0 < \theta' + 1/2$. Let p be a given number such that $0 < p < 1$. Take $y_0 = (1-p)\theta_0 + p(\theta' - 1/2)$, for $\theta_0 < \theta'$ and $y_0 = (1-p)\theta_0 + p(\theta' + 1/2)$, for $\theta' < \theta_0$.

(Note that when $p=1$, we get the first case.) Then, k_2 is taken as follows:

$$(12) \quad k_2 = \begin{cases} -k_1((1-p)\theta_0 + p(\theta' - 1/2)), & \text{for } \theta_0 < \theta' \\ -k_1((1-p)\theta_0 + p(\theta' + 1/2)), & \text{for } \theta' < \theta_0. \end{cases}$$

Substituting these values into $z(y)$ we have

$$(13) \quad z(y) = \begin{cases} [k_1\{y - ((1-p)\theta_0 + p(\theta' - 1/2))\}]^{1/(n-1)}, & \text{for } \theta_0 < \theta', \\ [k_1\{y - ((1-p)\theta_0 + p(\theta' + 1/2))\}]^{1/(n-1)}, & \text{for } \theta' < \theta_0. \end{cases}$$

which is drawn by stripe lines in the figure, below.

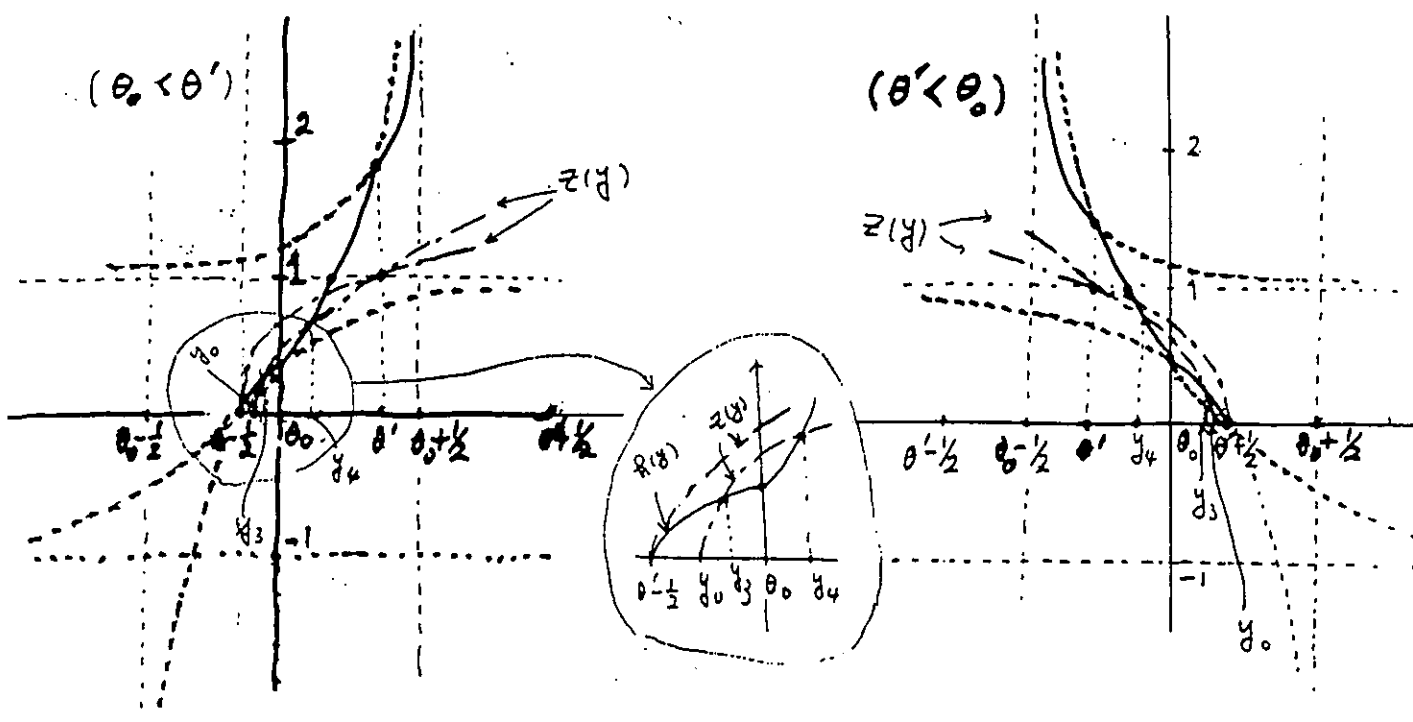
As before, since we need $h(\theta_0) < z(\theta_0)$, we take $k_1 = 2/p$ for $\theta_0 < \theta'$ and $k_1 = -2/p$ for $\theta' < \theta_0$. Then, $(z(\theta_0))^{n-1} = h(\theta_0)$, for all n and hence, there exists an integer n_0 such that for all $n > n_0$, $h(\theta_0) < z(\theta_0) < 1$.

Since $z(y) \geq 0$ and $dz(y)/dy > 0$ for all $y \geq y_0$, we get two intersection points of $z(y)$ and $h(y)$ for $\theta' < \theta_0$ and for $\theta_0 < \theta'$, respectively. Let y_3 and y_4 be such y -coordinates of those intersection points. Then, we finally have the UMP-test of form

$$(14) \quad \phi(y) = \begin{cases} 1, & \text{if } y < y_3 \text{ or } y > y_4 \\ 0, & \text{if } y_3 < y < y_4. \end{cases}$$

If we take $y_3 = \theta_0 - r$ and $y_4 = \theta_0 + r$ ($r = (1 - \alpha^{1/n})/2$), we get ϕ^* in (5) in D. P. No. 713. This is unbiased size- α from its construction. Hence, ϕ^* is UMP-unbiased size- α . (q.e.d.)

FIGURE. The Graph of $h(y)$.



Note that above result is immediately generalized to the case that the random sample of X_1, \dots, X_n is taken from the density $f(x|\theta) = c_0^{-1} I_{(\delta_1, \delta_2)}(y-\theta)$ ($\delta_1 < \delta_2$) where $c_0 = \delta_2 - \delta_1$. Let $Y = 2^{-1}(X_{(1)} + X_{(n)}) - \theta_0$ ($\theta_0 = (\delta_1 + \delta_2)/2$). Then, the p.d.f. of Y is given of form

$$(15) \quad g(y|\theta) = nc_0^{-n} \{c_0 - 2|y - \theta|\}^{n-1}, \quad \text{for } -c_0/2 < y - \theta < c_0/2.$$

Hence, (6) becomes

$$(16) \quad h(y) = (c_0 - 2|y - \theta'|) / (c_0 - 2|y - \theta_0|) \geq (k_2 + k_1 y)^{1/(n-1)},$$

for $\theta_0 < \theta'$ & $\theta' - c_0/2 < y < \theta_0 + c_0/2$ or for $\theta' < \theta_0$ & $\theta_0 - c_0/2 < y < \theta' + c_0/2$,

so on.

So, we can proceed the proof in the same way as above and get the UMP-test of form (14).

Remark. The author is extending the results of D.P.'s No.507 (1992), No. 627 (1995), No.713 (1997) and this paper. Those will come up soon.